

Statistical Kernel Estimators for Data Analysis and Exploration Tasks – Theory and Applications

PIOTR KULCZYCKI *

Center for Statistical Data Analysis Methods (*Head*)

Systems Research Institute

Polish Academy of Sciences

ul. Newelska 6, PL-01-447 Warsaw

POLAND

kulczycki@ibspan.waw.pl <http://www.ibspan.waw.pl/kulczycki>

Abstract: - At present, statistical kernel estimators constitute the dominant – in practice – method of nonparametric estimation. It allows the useful characterization of probability distributions without arbitrary assumptions regarding their membership to a fixed class. In this paper their use to the basic tasks of data analysis and exploration, i.e. identification of outliers, clustering, and classification, will be considered. In every case the final result will be an algorithm ensuring that its practical implementation does not demand of the user detailed knowledge of the theoretical aspects, or laborious research and calculations. The above presented theory has been successfully applied to various practical problems of engineering and management. Two of these, the design of a fault detection and diagnosis system for automatic control purposes, and a marketing support strategy for a mobile phone operator, will be demonstrated in detail. Useful procedures for the reduction of dimensionality and size of a random sample, subordinated to the specificity of kernel estimators, will also be commented.

Key-Words: - Information technology, Data analysis and exploration, Nonparametric estimation, Kernel estimators, Automatic control, Fault detection, Decision support, Marketing strategy

1 Introduction

Thanks to the dynamic development in contemporary computer systems, the range for practical application of nonparametric methods for identification and estimation is constantly growing. While the classical parametrical procedures arbitrarily assume the form of the function under investigations and then specify its parameters, the nonparametric methods do not require any such kind of limiting assumption.

The subject of this paper, currently dominant among nonparametric methods, is kernel estimators, mainly used to identify the most universal characteristic of a random variable – its distribution density. Here also is presented the application of this type of estimators in some basic tasks of data analysis and exploration – recognition of atypical elements (outliers), clustering, and classification – used next as examples in fault detection and diagnosis of industrial devices working in a real-time regime, and then to define a marketing support strategy for a mobile phone operator.

The following text also contains results of research in the field of kernel estimators carried out together with

M. Charytanowicz, K. Daniel, P.A. Kowalski, S. Lukasik, A. Mazgaj, C. Prochot, J. Waglowski, and S. Zak. This material was presented more broadly in the survey works [7, 8, 10, 12].

2 Kernel Estimators

Let the n -dimensional random variable $X : \Omega \rightarrow \mathbb{R}^n$, with a distribution having the density f , be given. Its kernel estimator $\hat{f} : \mathbb{R}^n \rightarrow [0, \infty)$ is calculated on the basis of the m -elements simple random sample x_1, x_2, \dots, x_m , experimentally obtained from the variable X , and is defined in its basic form by the formula

$$\hat{f}(x) = \frac{1}{mh^n} \sum_{i=1}^m K\left(\frac{x-x_i}{h}\right), \quad (1)$$

where the measurable, symmetrical with respect to zero and having a weak global maximum in this point, function $K : \mathbb{R}^n \rightarrow [0, \infty)$ fulfils the condition $\int_{\mathbb{R}^n} K(x) dx = 1$ and is called a kernel, whereas the positive coefficient h is referred to as a smoothing parameter. It is worth noting that a kernel estimator allows the identification of density for practically every distribution, without any assumptions concerning its

* Also: Cracow University of Technology, Department of Automatic Control (*Head*), ul. Warszawska 24, PL-31-155 Cracow, Poland, e-mail: kulczycki@pk.edu.pl, <http://www.control.pk.edu.pl/kulczycki>.

membership to a fixed class. Atypical, complex distributions, also multimodal, are regarded here as textbook unimodal. In the multidimensional case this enables, among others, the discovery of total dependences between particular coordinates of the random variable under investigation.

Setting of the quantities introduced in definition (1), i.e. choice of the form of the kernel K as well as calculation of the value for the smoothing parameter h , is most often carried out according to the criterion of minimum of an integrated mean-square error. Broader discussion and practical algorithms are found in the books [8; 22, 23]¹. In particular, the choice of the kernel form has no practical meaning and thanks to this it is possible to take into account primarily properties of the estimator obtained (e.g. its class of regularity, boundary of a support, etc.) or aspects of calculations, advantageous from the point of view of the applicational problem under consideration. Practical applications may also use additional procedures, some generally improving the quality of the estimator, and others – optional – possibly fitting the model to an existing reality. For the first group one should recommend the modification of the smoothing parameter [8 – Section 3.1.6; 22 – Section 5.3.1] and a linear transformation [8 – Section 3.1.4; 22 – Section 4.2.1], while for the second, the boundaries of a support [8 – Section 3.1.8; 22 – Section 2.10].

Kernel estimators allow effective modeling of the distribution density – a basic functional characteristic of random variables. Consequently this is fundamental in obtaining other functional characteristics and parameters. For example, if in a one-dimensional case the kernel K is such chosen that its primitive $I(x) = \int_{-\infty}^x K(y) dy$ may be analytically obtained, then the estimator of the distribution function

$$\hat{F}(x) = \frac{1}{m} \sum_{i=1}^m I\left(\frac{x-x_i}{h}\right) \quad (2)$$

can be easily calculated. Next, if the kernel K has (strictly) positive values, the solution for the equation

$$\hat{F}(x) = r \quad (3)$$

constitutes the kernel estimator of quantile of the order $r \in (0,1)$. For details and proofs of strong consistencies see [18].

¹ For calculating a smoothing parameter one can especially recommend the plug-in method in the one-dimensional case [8 – Section 3.1.5; 22 – Section 3.6.1], as well as the cross-validation method [8 – Section 3.1.5; 21 – Section 3.4.3] in the multidimensional. Comments for the choice of kernel may best be found in [8 – Section 3.1.3; 22 – Sections 2.7 and 4.5].

It is worth mentioning also the possibility of applying data compensation and dimensionality reduction procedures – original and useful algorithms can be found in the book [21 – Sections 2.5 and 3.4]. A specialized algorithm, based on simulated annealing, and dedicated to data analysis and exploration procedures found in the next chapter, is being researched – preliminary results are presented in the work [19].

3 Data Analysis and Exploration

The application of kernel estimators for recognition of atypical elements, clustering, and classification will be subsequently investigated in further sections of this chapter. In all three cases the n -dimensional random variable $X : \Omega \rightarrow \mathbb{R}^n$ is considered.

The text below also contains material from research carried out together with M. Charytanowicz, K. Daniel, and C. Prochot, published in the common works [14-17, 20].

3.1 Recognition of Atypical Elements

In many problems of data analysis, the task of recognizing atypical elements (outliers) – those which differ greatly from the general population – arises. This enables the elimination of such elements from the available set of data, which increases its homogeneity (uniformity), and facilitates analysis, especially in complex and unusual cases. In practice, the recognition process for outliers is most often carried out using procedures of statistical hypothesis testing [2]. The significance test based on the kernel estimators methodology will now be described.

Let therefore the random sample x_1, x_2, \dots, x_m treated as representative, therefore including a set of elements as typical as possible, be given. Furthermore, let $r \in (0,1)$ denote an assumed significance level. The hypothesis that $\tilde{x} \in \mathbb{R}^n$ is a typical element will be tested against the hypothesis that it is not, and therefore should be treated as an outlier. The statistic $S : \mathbb{R}^n \rightarrow [0, \infty)$, used here, can be defined by

$$S(\tilde{x}) = \hat{f}(\tilde{x}) \quad , \quad (4)$$

where \hat{f} denotes a kernel estimator of density, obtained for the random sample x_1, x_2, \dots, x_m mentioned above, while the critical set takes the left-sided form $(-\infty, a]$, when a constitutes the kernel estimator of quantile of the order r (see the text connected with formula (3)), calculated for the sample $\hat{f}(x_1), \hat{f}(x_2), \dots, \hat{f}(x_m)$, with the assumption that random variable support is

bounded to nonnegative numbers.

3.2 Clustering

The aim of clustering is the division of a data set – for example given in the form of the random sample x_1, x_2, \dots, x_m – into subgroups (clusters), with every one including elements “similar” to each other, but with significant differences between particular subgroups [5]. In practice this often allows the decomposition of a large data set with differing characteristics of elements into subsets containing elements of similar properties, which considerably facilitates further analysis, or even makes it possible at all. The following clustering procedure based on kernel estimators, taking advantage of the gradient methods concept [4] will be presented in this section.

Here the natural assumption is made that clusters are associated to modes – local maximums of the density kernel estimator \hat{f} , calculated for the considered random sample x_1, x_2, \dots, x_m . Within this procedure, particular elements are moved in a direction defined by a gradient, according to the following iterative algorithm:

$$x_j^0 = x_j \quad \text{for } j=1, 2, \dots, m \quad (5)$$

$$x_j^{k+1} = x_j^k + b \frac{\nabla \hat{f}(x_j^k)}{\hat{f}(x_j^k)} \quad \text{for } j=1, 2, \dots, m \text{ and } k=0, 1, \dots, \quad (6)$$

where $b > 0$ and ∇ denotes a gradient. Thanks to the proper choice of form of the kernel K , a suitable analytical formula for the gradient $\nabla \hat{f}$ becomes possible. As a result of the following iterative steps, the elements of the random sample move successively, focusing more and more clearly on a certain number of clusters. They can be defined after completing the k^* -th step, where k^* means the smallest number k such that $|D_k - D_{k-1}| \leq c D_0$, where $c > 0$ and $D_0 = \sum_{i=1}^m \sum_{j=i+1}^m d(x_i, x_j)$, $D_{k-1} = \sum_{i=1}^m \sum_{j=i+1}^m d(x_i^{k-1}, x_j^{k-1})$, $D_k = \sum_{i=1}^m \sum_{j=i+1}^m d(x_i^k, x_j^k)$, i.e. they are the sums of the distances between particular elements of the random sample under consideration before the beginning of algorithm (5)-(6) and having performed the $(k-1)$ -th and k -th steps, respectively.

Thus, after k^* -th step, one should calculate the kernel estimator for mutual distances of the elements $x_1^{k^*}, x_2^{k^*}, \dots, x_m^{k^*}$ (under the assumption of nonnegative support of the random variable), and next, the value can

be found where this estimator takes on the local minimum for the smallest value of its argument, omitting a possible minimum in zero. Finally, particular clusters are assigned those elements, whose distance to at least one of the others, is not greater than the above value. Thanks to the possibility of change in the smoothing parameter value, it becomes possible to affect the range of a number of obtained clusters, albeit without arbitrary assumptions concerning the strict value of this number, which allows it to be suited to a true data structure. Moreover, possible changes in the intensity of the smoothing parameter modification procedure enable influence on the proportion of clusters located in dense areas of random sample elements to the number of clusters on the “tails” of the distribution.

The detailed description of the above procedure can be found in the papers [14, 15].

3.3 Classification

The application of kernel estimators in a classification task will be considered now. Let the number $J \in \mathbb{N} \setminus \{0,1\}$ be given. Assume also, that the possessed random sample x_1, x_2, \dots, x_m has been divided into $J \in \mathbb{N} \setminus \{0,1\}$ nonempty and separate subsets

$$x_1^1, x_2^1, \dots, x_{m_1}^1 \quad (7)$$

$$x_1^2, x_2^2, \dots, x_{m_2}^2 \quad (8)$$

⋮

$$x_1^J, x_2^J, \dots, x_{m_J}^J, \quad (9)$$

while $\sum_{j=1}^J m_j = m$, representing J classes with features as mutually different as possible. The classification task requires deciding into which of them the given element $\tilde{x} \in \mathbb{R}^n$ should be reckoned [5].

The kernel estimators methodology provides a natural mathematical tool for solving the above problem in the optimal – in the sense of minimum for expectation of losses – Bayes approach. Let thus $\hat{f}_1, \hat{f}_2, \dots, \hat{f}_J$ denote kernel estimators of density calculated for subsets (7)-(9), respectively, treated here as samples. If sizes m_1, m_2, \dots, m_J are proportional to the “frequency” of appearance of elements from particular classes, the considered element \tilde{x} should be reckoned into the class for which the value

$$m_1 \hat{f}_1(\tilde{x}), m_2 \hat{f}_2(\tilde{x}), \dots, m_J \hat{f}_J(\tilde{x}) \quad (10)$$

is the greatest.

4 Fault Detection

The fault detection and diagnosis problem has lately become one of the most important challenges in modern control engineering. Early discovery of anomalies appearing in the operation of a controlled system, from an industrial robot to a nuclear reactor, most often allows serious incidents and even catastrophes to be avoided, which could save material damage, or loss of human life. Secondly, confirmation of kind and location of these anomalies is of fundamental meaning, especially when supervising large systems like complex chemical installations, as well as modern ships and airplanes. The importance of the above actions is multiplied by a psychological factor expressed by an increased feeling of safety, as well as – for the producer – prestige and commercial reputation. Finally, economic reasons often translate into a significant decrease in running costs, above all by ensuring the proper technological conditions as well as rationalizing overhauls and reducing repairs. Among the many different procedures used with this aim, the most universal are statistical methods. This paper presents the concept of a fault detection system, based on the kernel estimators methodology, covering:

- detection, so discovery of the existence of potential anomalies in the technical state of a device under supervision;
- diagnosis, that is identification of these anomalies;
- prognosis, i.e. warning of the threat of their occurrence in the near future, together with anticipated classification.

The procedures presented in Chapter 3 provide a complete and methodologically consistent mathematical tool to design an effective fault detection system for dynamical systems, covering detection, diagnosis, and also prognosis associated with them.

Assume that the technical state of a device under supervision may be characterized by a finite number of quantities measurable in real-time. These will be denoted in the form of the vector $x \in \mathbb{R}^n$, called a symptom vector. One can interpret this name noting that symptoms of any occurring anomalies should find the appropriate reflection in the features of a such-defined vector. More strictly, it is required that both correct functioning conditions and any type of diagnosed fault are connected with the most different sets of values and/or dissimilar relations between coordinates of the above vector as possible.

Assume also the availability of a fixed set of values of the symptom vector, representative for correct functioning conditions of a supervised device:

$$x_1, x_2, \dots, x_{m_0} , \tag{11}$$

as well as the set

$$x_1, x_2, \dots, x_M , \tag{12}$$

characteristic in the case of occurrence of anomalies. From the point of view of transparency of the designed fault detection system, in particular its function of diagnosis, it is worth dividing set (12) into $J \in \mathbb{N} \setminus \{0,1\}$ the most possibly different – in the sense of the values of particular coordinates of the symptom vector and/or relations between them – subsets assigned to the previously assumed types of diagnosed faults:

$$x_1^1, x_2^1, \dots, x_{m_1}^1 \tag{13}$$

$$x_1^2, x_2^2, \dots, x_{m_2}^2 \tag{14}$$

⋮

$$x_1^J, x_2^J, \dots, x_{m_J}^J , \tag{15}$$

while $\sum_{j=1}^J m_j = M$. Where there is no such division, one can automatically divide set (12) into subsets (13)-(15) using the clustering algorithm presented in Section 3.2, although this then often requires laborious interpretation concerning each of them.

Fault detection will first be considered. With this aim the procedure for the recognition of atypical elements, described in Section 3.1, can be applied. Assume therefore that the random sample considered there, including elements treated as typical, constitutes set (11) representing the correct functioning conditions for a supervised device, while \tilde{x} denotes its current state. Applying the above mentioned procedure for the recognition of atypical elements, one can confirm if the present conditions should be regarded as typical or rather not, thus showing the appearance of anomalies.

For fault diagnosis, if one already is in possession of samples (13)-(15) characterizing particular types of faults being diagnosed, then after the above described detection of anomalies, one can – applying directly the procedure for Bayes classification presented in Section 3.3 – infer which of them is being dealt with. Note that the range of faults which can be discovered by detection may significantly exceed all types of faults assumed to be diagnosed.

Finally, if subsequent values of the symptom vector, obtained successively during the supervising process, are available, then it is possible to realize fault prognosis. It can be carried out by separate forecasts of values of the function \hat{f} given by dependence (4) and $m_1 \hat{f}_1, m_2 \hat{f}_2, \dots, m_J \hat{f}_J$ to be seen in formula (10), and inferences based on these forecasts for detection and

diagnosis, according to guidelines presented in the previous two paragraphs. To calculate the values of forecasts of the functions \hat{f} , \hat{f}_1 , $\hat{f}_2, \dots, \hat{f}_J$ it is recommended to use the classical linear regression method separately, though in a version enabling easy updating of a model during successive collection of subsequent current values of the symptom vector. Appropriate formulas are found in the book [1 – Chapter 3 and additionally Chapter 4].

The proper operation of the fault detection system investigated in this section was verified experimentally for a robust control applied to the task from a field of robotics [9]. Thus, in cases where the symptoms appeared abruptly, the anomalies of the device were promptly discovered and correctly recognized within the scope of detection and diagnosis. If, on the other hand, the fault was accompanied by a slow progression of symptoms, it was forecast with a correct indication of the type of fault about to occur (scope of prognosis), and later it was also discovered and identified in detection and diagnosis. One should underline that fault prognosis, still rare in practical applications, proved to be highly effective in the case of slowly progressing symptoms, discovering and identifying anomalies before the object's characteristics transgressed the range for correct conditions for a system's functioning, thanks to the proper recognition of the change in the trend of values of the symptom vector, which indicates an unfortunate direction of its evolution.

More details can be found in the paper [13].

5 Marketing Strategy

The highly dynamic growth prevalent on the mobile phone network market, naturally necessitates a company to permanently direct its strategy towards satisfying the differing needs of its clients, while at the same time maximizing its income. The uncontrollable nature of this kind of activity, however, can lead to a loss of coherence in treating particular clients, and their subsequent defection to competitors. To avoid this a formal solution of global nature must be found. Below are presented results of research obtained using statistical kernel estimators, carried out in the procedures discussed in Chapter 3. This procedure, prepared for a Polish mobile phone network operator, concerns long term business clients, i.e. those with more than 30 SIM cards and an account history of at least 2 years.

In practice there is a vast spectrum of quantities characterizing particular subscribers. Following detailed analysis of the economic aspects of the task under investigation here, it was taken that basic traits of clients would be shown by three quantities: average monthly income per SIM card, length of subscription and number

of active SIM cards. Thus each of m -elements of a database x_1, x_2, \dots, x_m is characterized by the following 3-dimensional vector:

$$x_i = \begin{bmatrix} x_{i,1} \\ x_{i,2} \\ x_{i,3} \end{bmatrix} \quad \text{for } i = 1, 2, \dots, m, \quad (16)$$

where $x_{i,1}$ denotes the average monthly income per SIM card of the i -th client, $x_{i,2}$ – its length of subscription, and $x_{i,3}$ – the number of active SIM cards.

In the initial phase, atypical elements of the set x_1, x_2, \dots, x_m (outliers) are eliminated, according to the procedure presented in Section 3.1. The uniformity of the data structure is so increased, and it is worth underling, this effect is obtained by canceling only those elements which would not be of importance further in the procedure investigated.

Next clustering of the data set is performed, using the algorithm as shown in Section 3.2. This results in a division of the data set representing specific clients, into groups of similar nature. This should be followed by another, albeit slightly different, elimination of atypical elements, achieved by removing clusters with small numbers of elements. As well as omitting information of little significance, it is also in order to improve conditions for the classification algorithm applied at a further stage of the procedure worked out – kernel estimators calculated on the basis of an insufficient number of samples (7)-(9) may not be representative.

Next for each of the above defined clusters, an optimal – from the point of view of expected profit of the operator – strategy is created for treating subscribers belonging to it. With regard to the imprecise evaluation of experts used here, elements of fuzzy logic [6] and preference theory [3] have been used – details are however beyond the scope of this paper.

It is worth pointing out that none of the above calculations must be carried out at the same time as negotiating with the client, but merely updated (in practice once every 1-6 months).

The client being negotiated with is described with the aid of three quantities, in reference to formula (16) given here as:

$$\tilde{x} = \begin{bmatrix} \tilde{x}_1 \\ \tilde{x}_2 \\ \tilde{x}_3 \end{bmatrix}, \quad (17)$$

where \tilde{x}_1 denotes its average monthly income per SIM card, \tilde{x}_2 – length of subscription, and \tilde{x}_3 is the number

of active SIM cards of that client. This data can relate to the subscriber history to date in a given network, when renegotiating contract terms, or in a rival network if attempting to take them over. Classification to the proper subscriber group, from those obtained as a result of clustering, is achieved with the procedure presented in Section 3.3. To this purpose, first kernel estimators are calculated separately for each cluster. In classification element (17) is mapped to the class (cluster), for which the value $m_1 \hat{f}_1(\tilde{x})$, $m_2 \hat{f}_2(\tilde{x}), \dots, m_J \hat{f}_J(\tilde{x})$ to be seen in formula (10), is greatest. Due to the fact that the marketing strategies for particular clusters have already been defined, this finally completes the procedure for the algorithm to support the marketing strategy for a business client, investigated here.

The above method was successfully implemented for the needs of a Polish network operator. In relation to the comment at the end of Section 3.2, the intensity of the smoothing parameter modification procedure was somewhat lessened with respect to the optimal in the mean-square sense, with the aim of dividing the largest cluster and decreasing the number of peripheral clusters. Finally, two large clusters containing 27% and 23% of elements, two medium of 14% and 7%, and 22 small each including less than 3% – the most uncommon firms, were obtained.

More details can be found in the papers [16, 17].

References:

- [1] Abraham B., Ledolter J., *Statistical Methods for Forecasting*, Wiley, 1983.
- [2] Barnett V., Lewis T., *Outliers in Statistical Data*, Wiley, 1994.
- [3] Fodor J., Roubens M., *Fuzzy Preference Modelling and Multicriteria Decision Support*, Kluwer, 1994.
- [4] Fukunaga K., Hostetler L.D., The estimation of the gradient of a density function, with applications in Pattern Recognition, *IEEE Transactions on Information Theory*, Vol. 21, 1975, pp. 32-40.
- [5] Hand J.H., Mannila H., Smyth P., *Principles of Data Mining*, MIT Press, 2001.
- [6] Kacprzyk J., *Zbiory rozmyte w analizie systemowej*, PWN, Warszawa, 1986.
- [7] Kulczycki P., Applicational Possibilities of Nonparametric Estimation of Distribution Density for Control Engineering, *Bulletin of the Polish Academy of Sciences; Technical Sciences*, Vol. 56, 2008, pp. 347-359.
- [8] Kulczycki P., *Estymatory jądrowe w analizie systemowej*, WNT, 2005.
- [9] Kulczycki P., Fuzzy Controller for Mechanical Systems, *IEEE Transactions on Fuzzy Systems*, Vol. 8, 2000, pp. 645-652.
- [10] Kulczycki P., Estymatory jądrowe w badaniach systemowych, in: *Techniki informacyjne w badaniach systemowych*, Kulczycki P., Hryniewicz O., Kacprzyk J. (eds.), WNT, 2007, pp. 79-105.
- [11] Kulczycki P., Kernel Estimators in Industrial Applications, in: *Soft Computing Applications in Industry*, Prasad B. (ed.), Springer-Verlag, 2008, pp. 69-91.
- [12] Kulczycki P., Nonparametric Estimation for Control Engineering, in: *Proc. 4th WSEAS/IASME International Conference on Dynamical Systems and Control*, Corfu, 26-28 October 2008, pp. 115-121.
- [13] Kulczycki P., Statistical Kernel Estimators for Design of a Fault Detection, Diagnosis and Prognosis System, *The Open Cybernetics and Systemics Journal*, Vol. 2, 2008, pp. 180-184, open access: <http://www.bentham.org/open/tocsj/openaccess2.htm>.
- [14] Kulczycki P., Charytanowicz M., A Complete Gradient Clustering Algorithm Formed with Kernel Estimators, *Applied Mathematics and Computer Science*, 2010, in press.
- [15] Kulczycki P., Charytanowicz M., Kompletny algorytm gradientowej klasteryzacji, in: *Sterowanie i automatyzacja: aktualne problemy i ich rozwiązania*, Malinowski K., Rutkowski L. (eds.), EXIT, 2008, pp. 312-321.
- [16] Kulczycki P., Daniel K., Algorytm wspomaganie strategii marketingowej operatora telefonii komórkowej, in: *Badania operacyjne i systemowe 2006: metody i techniki*, Kacprzyk J., Budzinski R. (eds.), EXIT, 2006, pp. 245-256.
- [17] Kulczycki P., Daniel K., Metoda wspomaganie strategii marketingowej operatora telefonii komórkowej, *Przegląd statystyczny*, 2009, in press.
- [18] Kulczycki P., Dawidowicz A.L., Kernel estimator of quantile, *Universitatis Iagellonicae Acta Mathematica*, Vol. XXXVII, 1999, pp. 325-336.
- [19] Kulczycki P., Lukasik S., Redukcja wymiaru i liczności próby dla potrzeb syntezy statystycznego układu wykrywania uszkodzeń, in: *Systemy wykrywające, analizujące i tolerujące usterki*, Kowalczyk Z. (ed.), PWNT, 2009, pp. 139-146.
- [20] Kulczycki P., Prochot C., Wykrywanie elementów odosobnionych za pomocą metod estymacji nieparametrycznej, in: *Badania operacyjne i systemowe: podejmowanie decyzji – podstawy teoretyczne i zastosowania*, Kulikowski R., Kacprzyk J., Slowinski R. (eds.), EXIT, 2004, pp. 313-328.
- [21] Pal S.K., Mitra P., *Pattern Recognition Algorithms for Data Mining*, Chapman and Hall, 2004.
- [22] Silverman B.W., *Density Estimation for Statistics and Data Analysis*, Chapman and Hall, 1986.
- [23] Wand M.P., Jones M.C., *Kernel Smoothing*. Chapman and Hall, 1994.