

# AN EM APPROACH FOR POISSON-GAUSSIAN NOISE MODELING

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2 September, 2011

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# THE IMPORTANCE OF THE POISSON-GAUSSIAN NOISE

## Where does it appear?

- ▶ CCD camera images [Healey *et al.* 1994]
- ▶ Medical images [Nichols *et al.* 2002]
- ▶ Biological images (fluorescence microscopy) [Pawley 1994]
- ▶ Astronomical images [Benvenuto *et al.* 2008]

## What are the signal processing challenges?

- ▶ noise parameters identification
- ▶ signal recovery

## When is it especially difficult?

- ▶ low level signal (The Anscombe transform introduces a significant bias)
- ▶ non-zero background (The model parameters cannot be estimated from first and second order statistics)

# GENERAL MODEL

Observations

$$\forall s \in \{1, \dots, S\}$$

$$\forall t \in \{1, \dots, T\}$$

$$R_{s,t} = \alpha Q_{s,t} + N_{s,t}$$

Assumptions:

$$1. (Q_{s,t})_{\substack{1 \leq s \leq S, \\ 1 \leq t \leq T}}$$

$$(N_{s,t})_{\substack{1 \leq s \leq S, \\ 1 \leq t \leq T}}$$

independent sequences

2.  $Q_{s,t}$  and  $N_{s,t}$  mutually independent

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Poisson noise

$$Q_{s,t} \sim \mathcal{P}(u_s)$$

$$(u_s)_{1 \leq s \leq S} \geq 0 -$$

“clean” image

$$\alpha \in ]0, +\infty[ -$$

scaling param.

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independent sequences

- $Q_{s,t}$  and  $N_{s,t}$  mutually  
 independent

Gaussian noise

$$N_{s,t} \sim \mathcal{N}(c, \sigma^2)$$

$c \in \mathbb{R}$  - mean  
 $\sigma^2$  - variance  
 $\sigma^2 \geq 0$

# GENERAL MODEL

Observations

$$\forall s \in \{1, \dots, S\}$$

$$\forall t \in \{1, \dots, T\}$$

$$R_{s,t} = \alpha Q_{s,t} + N_{s,t}$$

Problem - to identify:

- ▶  $\alpha$
- ▶  $c$
- ▶  $\sigma^2$

$u_s$  - unknown

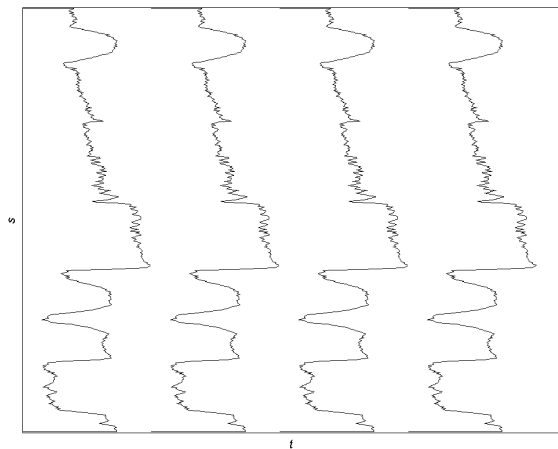
Poisson noise

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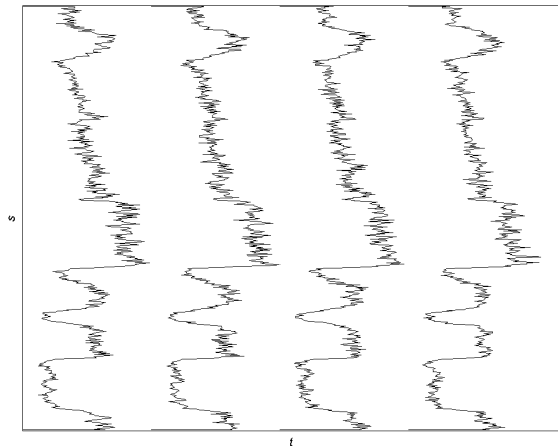
## EXAMPLE

 $U_{s,t}$  $T = 4,$   
 $S = 512$ 

Original signal

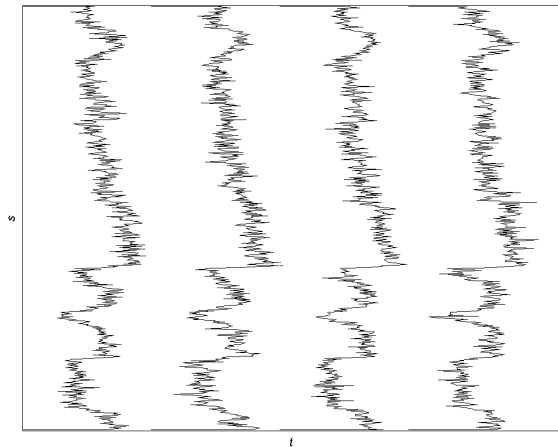


## EXAMPLE

 $Q_{s,t}$  $T = 4,$   
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Signal corrupted by Poisson noise

## EXAMPLE

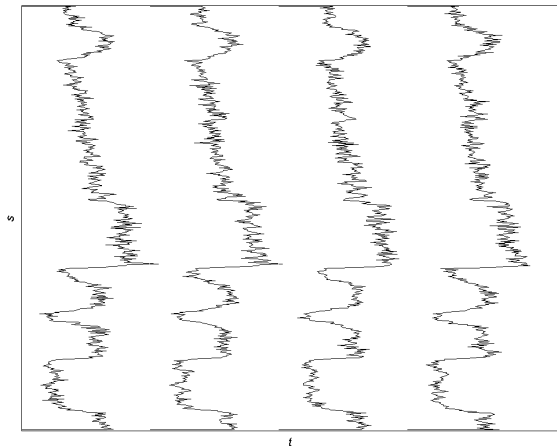


$$U_{s,t} + N_{s,t}$$

$$T = 4,$$
$$S = 512$$

Signal corrupted by Gaussian noise

## EXAMPLE



$$\alpha Q_{s,t} + N_{s,t}$$

$$T = 4,$$
$$S = 512$$

Signal corrupted by Poisson and Gaussian noises


# MOMENT BASED APPROACH

$$\kappa_n[R_{s,t}] = \alpha^n \kappa_n[Q_{s,t}] + \kappa_n[N_{s,t}]$$

# MOMENT BASED APPROACH

$$\kappa_n[R_{s,t}] = \alpha^n \kappa_n[Q_{s,t}] + \kappa_n[N_{s,t}]$$

Cumulant of order  $n$



# MOMENT BASED APPROACH

$$\kappa_n[R_{s,t}] = \alpha^n \kappa_n[Q_{s,t}] + \kappa_n[N_{s,t}]$$

Then:

- ▶ mean value

$$\kappa_1[R_{s,t}] = \mathbf{E}[R_{s,t}] = \alpha u_s + c$$

- ▶ variance

$$\kappa_2[R_{s,t}] = \mathbf{Var}[R_{s,t}] = \alpha^2 u_s + \sigma^2$$

- ▶ higher-order cumulants

$$\kappa_n[R_{s,t}] = \alpha^n u_s, \quad n \geq 3$$

# PROPOSED CUMULANT BASED METHOD

From:

$$\begin{cases} \mathbf{E}[R_{s,t}] = \alpha u_s + c \\ \mathbf{Var}[R_{s,t}] = \alpha^2 u_s + \sigma^2 \end{cases}$$

After some algebraic manipulations, we obtain:

$$\hat{\alpha} = \frac{S \sum_{s=1}^S \hat{\mathbf{E}}[r_{s,t}] \widehat{\mathbf{Var}}[r_{s,t}] - \sum_{s=1}^S \hat{\mathbf{E}}[r_{s,t}] \sum_{s=1}^S \widehat{\mathbf{Var}}[r_{s,t}]}{S \sum_{s=1}^S (\hat{\mathbf{E}}[r_{s,t}])^2 - (\sum_{s=1}^S \hat{\mathbf{E}}[r_{s,t}])^2}$$

- The estimation of  $\hat{\alpha}$  is precise as only cumulants of order 1 and 2 are used

# PROPOSED CUMULANT BASED METHOD

Estimate of  $\sigma^2$ :

$$\hat{\sigma}^2 = \sum_{s \in \mathbb{I}} \widehat{\text{Var}}[r_{s,t}] - \hat{\alpha}^{-1} \hat{\kappa}_3[r_{s,t}]$$

Estimation error increases:

- ▶ when  $T$  is small
- ▶ when  $u_s$  is large



# PROPOSED CUMULANT BASED METHOD

Estimate of  $\sigma^2$ :

$$\hat{\sigma}^2 = \frac{\sum_{s \in \mathbb{I}} \widehat{\text{Var}}[r_{s,t}]^{-6} (\widehat{\text{Var}}[r_{s,t}] - \hat{\alpha}^{-1} \hat{\kappa}_3[r_{s,t}])}{\sum_{s \in \mathbb{I}} \widehat{\text{Var}}[r_{s,t}]^{-6}}$$

Weighted  
estimator

# PROPOSED CUMULANT BASED METHOD

Estimate of  $\sigma^2$ :

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Weighted  
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Set  $\mathbb{I}$  may be defined as:

$$\mathbb{I} = \{s \in \{1, \dots, S\} \mid \widehat{\text{Var}}[r_{s,t}] - \hat{\alpha}^{-1} \hat{\kappa}_3[r_{s,t}] \geq 0\}$$

- Estimate  $\hat{\sigma}^2$  is precise, because weighting the observations effectively discards the unreliable ones.

# PROPOSED CUMULANT BASED METHOD

Estimate of  $c$ :

$$\hat{c} = \frac{1}{S} \sum_{s=1}^S \left( \widehat{\mathbf{E}}[r_{s,t}] - \hat{\alpha}^{-1} \widehat{\mathbf{Var}}[r_{s,t}] \right) + \frac{\hat{\sigma}^2}{\hat{\alpha}}$$

Summary:

- ▶ All parameters estimated using cumulants of order at most 3
- ▶  $\alpha$  estimator is a function of 1<sup>st</sup> and 2<sup>st</sup> order moments
- ▶ Precise estimator of  $\sigma^2$
- ▶  $c$  derived from  $\alpha$  and  $\sigma^2$

## RESULTS

S	T	$\hat{\sigma}^2$		$\hat{c}$		$\hat{\alpha}$		$\overline{\text{err}}$
		bias	std	bias	std	bias	std	
1024	50	-14.83	20.29	-1.74	5.07	-0.21	0.14	0.22
	100	-8.58	13.86	-0.54	4.04	-0.09	0.11	0.16
	200	-11.69	10.13	1.18	2.7	-0.04	0.04	0.14
	500	-1.93	6.24	-0.26	1.61	-0.02	0.04	0.06
	1000	-1.16	4.55	-0.24	1.1	-0.01	0.03	0.04

$\sigma^2$	$\hat{\sigma}^2$		$\hat{c}$		$\hat{\alpha}$		$\overline{\text{err}}$
	bias	std	bias	std	bias	std	
0.25	-0.14	0.03	-0.11	5.15	-0.19	0.13	0.18
25	-3.59	3.89	-0.18	4.19	-0.19	1.14	0.19
400	-11.45	84.32	-0.84	10.53	-0.22	0.15	0.34

c	$\hat{\sigma}^2$		$\hat{c}$		$\hat{\alpha}$		$\overline{\text{err}}$
	bias	std	bias	std	bias	std	
5	-11.70	19.43	-1.20	5.07	-0.20	0.13	0.33
15	-10.16	26.15	-0.72	5.23	-0.18	0.13	0.17
100	-11.47	24.10	0.21	5.19	-0.17	0.13	0.09

$\alpha$	$\hat{\sigma}^2$		$\hat{c}$		$\hat{\alpha}$		$\overline{\text{err}}$
	bias	std	bias	std	bias	std	
1	12.99	6.17	13.73	6.29	-0.02	0.02	0.64
5	-2.93	5.69	-0.25	2.83	-0.09	0.07	0.15
50	1.07	56.4	-3.55	25.7	-1.10	0.67	0.81

The reliability of cumulant based method:

- ▶ increases with  $T$
- ▶ decreases as  $\sigma^2$  and  $\alpha$  increase
- ▶ decreases for small value of  $\alpha$  and  $c$

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# MAXIMUM LIKELIHOOD ESTIMATOR

$$(\hat{u}, \hat{\alpha}, \hat{c}, \hat{\sigma}) = \underset{(u, \alpha, c, \sigma)}{\operatorname{argmax}} f_R(r | u, \alpha, c, \sigma)$$

Joint  
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for all ob-  
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 $r$

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$$f_R(r | u, \alpha, c, \sigma) = \sum_{q \in \mathcal{N}^{ST}} p_{R,Q}(r, q | u, \alpha, c, \sigma)$$

Mixed  
continuous-  
discrete  
probability  
distribu-  
tion of  
 $(R, Q)$

# MAXIMUM LIKELIHOOD ESTIMATOR

$$(\hat{u}, \hat{\alpha}, \hat{c}, \hat{\sigma}) = \operatorname{argmax}_{(u, \alpha, c, \sigma)} f_R(r | u, \alpha, c, \sigma)$$

Joint density function for all observations  $r$

$$f_R(r | u, \alpha, c, \sigma) = \sum_{q \in \mathcal{N}^{ST}} p_{R,Q}(r, q | u, \alpha, c, \sigma)$$

Mixed continuous-discrete probability distribution of  $(R, Q)$

$$p_{R,Q}(r, q | u, \alpha, c, \sigma) =$$



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( $R, Q$ )

$$p_{R,Q}(r, q | u, \alpha, c, \sigma) = f_{R|Q=q}(r | u, \alpha, c, \sigma) \mathbf{P}(Q = q | u)$$

# MAXIMUM LIKELIHOOD ESTIMATOR

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Joint density function for all observations  $r$

$$f_R(r | u, \alpha, c, \sigma) = \sum_{q \in \mathbb{N}^{ST}} p_{R,Q}(r, q | u, \alpha, c, \sigma)$$

Mixed continuous-discrete probability distribution of  $(R, Q)$

$$p_{R,Q}(r, q | u, \alpha, c, \sigma) = f_N(r - \alpha q | c, \sigma) \mathbf{P}(Q = q | u)$$

# MAXIMUM LIKELIHOOD ESTIMATOR

Joint density function  $f_R(r | u, \alpha, c, \sigma)$  for Poisson + Gaussian:

$$\frac{1}{(2\pi)^{ST/2} \sigma^{ST}} \prod_{s=1}^S \exp(-Tu_s) \prod_{t=1}^T \sum_{q_{s,t}=1}^{+\infty} \exp\left(-\frac{(r_{s,t} - \alpha q_{s,t} - c)^2}{2\sigma^2}\right) \frac{u_s^{q_{s,t}}}{q_{s,t}!}$$

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The likelihood takes an intricate form which renders the computation of the ML estimator too difficult.

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Expectation-Maximization algorithm

# EXPECTATION-MAXIMIZATION ALGORITHM

## Expectation step

$$J(\theta \mid \theta^{(n)}) = \mathbf{E}_{Q \mid R=r, \theta^{(n)}} [\ln p_{R,Q}(\text{R}, \text{Q} \mid \theta)]$$

## Maximization step

$$(\forall n \in \mathbb{N}) \quad \theta^{(n+1)} = \underset{\theta}{\operatorname{argmin}} - J(\theta \mid \theta^{(n)})$$

## Importance of having a good enough initialization

Proposed cumulant-based method

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Observed data

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Observed data

Vector of parameters:

## Maximization step

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Observed data

Vector of parameters:

$$\blacktriangleright \theta = (u, \alpha, c, \sigma^2)$$

## Maximization step

$$(\forall n \in \mathbb{N}) \quad \theta^{(n+1)} = \underset{\theta}{\operatorname{argmin}} - J(\theta \mid \theta^{(n)})$$

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# EXPECTATION-MAXIMIZATION ALGORITHM

## Expectation step

$$J(\theta \mid \theta^{(n)}) = \mathbf{E}_{Q \mid R=r, \theta^{(n)}} [\ln p_{R,Q}(R, Q \mid \theta)]$$

Observed data

Vector of parameters:

$$\blacktriangleright \theta = (u, \alpha, c, \sigma^2)$$

Missing data

## Maximization step

$$(\forall n \in \mathbb{N}) \quad \theta^{(n+1)} = \underset{\theta}{\operatorname{argmin}} - J(\theta \mid \theta^{(n)})$$

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## EM MAXIMIZATION STEP

$$(\forall s \in \{1, \dots, S\}) \quad u_s^{(n+1)} = \frac{1}{T} \sum_{t=1}^T \mathbf{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}]$$

$$\begin{aligned} & \begin{bmatrix} ST & \mathbf{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}] \\ \mathbf{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}] & \sum_{s,t} \mathbf{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}^2] \end{bmatrix} \begin{bmatrix} c^{(n+1)} \\ \alpha^{(n+1)} \end{bmatrix} \\ &= \begin{bmatrix} \sum_{s,t} r_{s,t} \\ \sum_{s,t} r_{s,t} \mathbf{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}] \end{bmatrix} \end{aligned}$$

$$(\sigma^2)^{(n+1)} =$$

$$\frac{1}{ST} \sum_{s=1}^S \sum_{t=1}^T r_{s,t} \left( r_{s,t} - \alpha^{(n+1)} \mathbf{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}] - c^{(n+1)} \right)$$

## EM MAXIMIZATION STEP

$$(\forall s \in \{1, \dots, S\}) \quad u_s^{(n+1)} = \frac{1}{T} \sum_{t=1}^T \mathbb{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}]$$

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$$= \begin{bmatrix} \sum_{s,t} r_{s,t} & \mathbb{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}] \\ \sum_{s,t} r_{s,t} \mathbb{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}] & \mathbb{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}] \end{bmatrix}$$

Updated in  
EM  
Expectation  
step

$$(\sigma^2)^{(n+1)} =$$

$$\frac{1}{ST} \sum_{s=1}^S \sum_{t=1}^T r_{s,t} \left( r_{s,t} - \alpha^{(n+1)} \mathbb{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}] - c^{(n+1)} \right)$$

# EM EXPECTATION STEP

$$\mathbb{E}_{Q|R=r, \theta^{(n)}}[Q_{s,t}] = \sum_{q_{s,t}=1}^{+\infty} q_{s,t} \mathbf{P}(Q_{s,t} = q_{s,t} | R = r, \theta^{(n)})$$

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After some calculations, we have:

$$\mathbb{E}_{Q|R=r, \theta^{(n)}}[Q_{s,t}] = \frac{\zeta_{s,t}^{(n)}}{\eta_{s,t}^{(n)}}$$



## EM EXPECTATION STEP

$$\mathbb{E}_{Q|R=r, \theta^{(n)}}[Q_{s,t}] = \sum_{q_{s,t}=1}^{+\infty} q_{s,t} \mathbf{P}(Q_{s,t} = q_{s,t} | R = r, \theta^{(n)})$$

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$$\mathbb{E}_{Q|R=r, \theta^{(n)}}[Q_{s,t}] = \frac{\zeta_{s,t}^{(n)}}{\eta_{s,t}^{(n)}} \sum_{q_{s,t}=1}^{+\infty} \exp\left(-\frac{(r_{s,t} - \alpha^{(n)}q_{s,t} - c^{(n)})^2}{2(\sigma^2)^{(n)}}\right) \frac{(u_s^{(n)})^{q_{s,t}}}{(q_{s,t} - 1)!}$$

## EM EXPECTATION STEP

$$\mathbb{E}_{Q|R=r, \theta^{(n)}}[Q_{s,t}] = \sum_{q_{s,t}=1}^{+\infty} q_{s,t} \mathbf{P}(Q_{s,t} = q_{s,t} | R = r, \theta^{(n)})$$

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Problem: Infinite sums

# EM EXPECTATION STEP

Similarly:

$$\mathbb{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}^2] = \sum_{q_{s,t}=1}^{+\infty} (q_{s,t})^2 \mathbb{P}(Q_{s,t} = q_{s,t} | R = r, \theta^{(n)})$$

After some calculations, we have:

$$\mathbb{E}_{Q|R=r, \theta^{(n)}} [Q_{s,t}^2] = \frac{\xi_{s,t}^{(n)}}{\eta_{s,t}^{(n)}} = \frac{\sum_{q_{s,t}=1}^{+\infty} q_{s,t} \exp\left(-\frac{(r_{s,t} - \alpha^{(n)} q_{s,t} - c^{(n)})^2}{2(\sigma^2)^{(n)}}\right) \frac{(u_s^{(n)})^{q_{s,t}}}{(q_{s,t} - 1)!}}{\sum_{q_{s,t}=0}^{+\infty} \exp\left(-\frac{(r_{s,t} - \alpha^{(n)} q_{s,t} - c^{(n)})^2}{2(\sigma^2)^{(n)}}\right) \frac{(u_s^{(n)})^{q_{s,t}}}{q_{s,t}!}}$$

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# EM EXPECTATION STEP

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**Problem:** Infinite sums

**Solution:** Adaptive truncation technique

# ALGORITHM

## Initialization:

$\theta^{(1)} \leftarrow (u^{(1)}, \alpha^{(1)}, c^{(1)}, (\sigma^2)^{(1)})$  (Cumulant Method)

## EM Algorithm:

for  $n \leftarrow 1$  to  $N$  do

### Expectation step

$$E_{Q|R=r, \theta^{(n)}}[Q_{s,t}] \leftarrow \frac{\zeta_{s,t}^{(n)}}{\eta_{s,t}^{(n)}}$$

$$E_{Q|R=r, \theta^{(n)}}[Q_{s,t}^2] \leftarrow \frac{\xi_{s,t}^{(n)}}{\eta_{s,t}^{(n)}}$$

### Maximization step

$\theta^{(n+1)} \leftarrow (u^{(n+1)}, \alpha^{(n+1)}, c^{(n+1)}, (\sigma^2)^{(n+1)})$

end for

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end for

# NUMERICAL RESULTS

Param.					Proposed EM algorithm						Error	
S	T	$\sigma^2$	c	$\alpha$	$\hat{\sigma}^2$		$\hat{c}$		$\hat{\alpha}$		$\overline{\text{err}}_{\text{cum}}$	$\overline{\text{err}}_{\text{EM}}$
					bias	std	bias	std	bias	std		
1024	50	100	10	10	-1.86	14.02	-0.35	1.79	-0.21	0.05	0.22	0.11
1024	100	100	10	10	0.54	10.92	0.27	1.38	-0.08	0.02	0.16	0.08
1024	200	100	10	10	-0.91	7.89	0.09	0.88	-0.04	0.03	0.14	0.05
1024	50	100	5	10	1.30	13.16	0.08	1.61	-0.20	0.05	0.33	0.15
1024	50	25	10	1	-0.28	0.50	0.00	0.48	-0.02	0.01	0.64	0.03
1024	50	400	10	10	7.51	45.24	0.71	5.13	-0.22	0.07	0.34	0.21

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- Our EM algorithm performs well even if  $T$  is small

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- ▶ Our EM algorithm performs well even if  $T$  is small
- ▶ Our EM does not appear to be sensitive to small values of  $\alpha$  and  $c$

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S	T	$\sigma^2$	c	$\alpha$	$\hat{\sigma}^2$		$\hat{c}$		$\hat{\alpha}$		$\overline{\text{err}}_{\text{cum}}$	$\overline{\text{err}}_{\text{EM}}$
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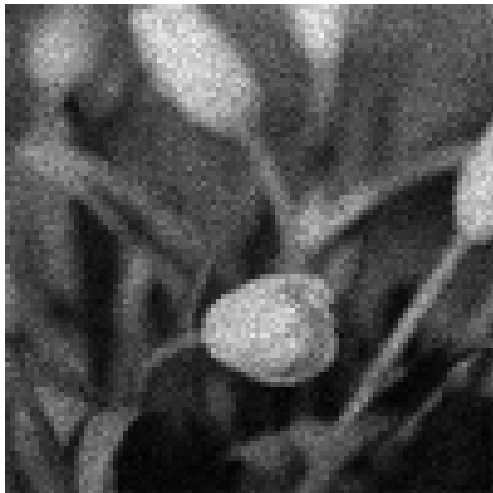
- ▶ Our EM algorithm performs well even if  $T$  is small
- ▶ Our EM does not appear to be sensitive to small values of  $\alpha$  and  $c$
- ▶ Our EM results improve when  $\sigma^2$  is high.

# VISUAL RESULTS



Original image. The range of values  $(0, 255)$ . The number of realizations  $T = 40$

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Original image. The range of values  $(0, 255)$ . The number of realizations  $T = 40$

Corrupted by noise with parameters  $\alpha = 50$ ,  $\sigma^2 = 416$  and  $c = 10$ .



# VISUAL RESULTS



Original image. The range of values  $(0, 255)$ . The number of realizations  $T = 40$

Corrupted by noise with parameters  $\alpha = 50$ ,  $\sigma^2 = 416$  and  $c = 10$ .

Reconstructed  $u_s$  values; SNR = 35.6 dB

# SUMMARY

- ▶ We have proposed a new EM-based approach dealing with Poisson plus Gaussian noise in case of non-zero background parameter estimation problems
- ▶ We have proposed a very accurate cumulant based initialization for this method
- ▶ The method is especially useful when a low level signal is expected
- ▶ These algorithms can estimate noise parameters for denoising or restoration procedures, in which they are often assumed to be known [Benvenuto 2008,Luisier 2011]
- ▶ We plan also to compare our method with a Bayesian approach